# Shannon Information and Entropy in the Analysis of Independent Components and Clusters 

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# Dependencies in Data 

# Shannon's Information and Entropy 

Independent Component Analysis

Clustering

Value of Information

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## Dependencies in data

- Consider the following records:

| Case: | Age | Gender | Income ( $£ \mathrm{~K})$ | Outcome <br> $\mathrm{K})$ | $(£$ | Home <br> owner |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Credit <br> score |  |  |  |  |  |  |
| 1 | 21 | 0 | 2 | 1 | 0 | 3 |
| 2 | 18 | 1 | 1 | 2 | 0 | 1 |
| 3 | 50 | 1 | 6 | 2 | 1 | 5 |
| 4 | 23 | 0 | 3 | 1 | 1 | 4 |
| 5 | 40 | 1 | 3 | 2 | 0 | 2 |

- Each case is a vector $y \in \mathbb{R}^{m}$ :

$$
\begin{aligned}
y^{1}= & (21,0,2,1,0,3)^{T} \\
y^{2}= & (18,1,1,2,0,1)^{T} \\
& \cdots \\
y^{n}= & (23,0,3,1,1,4)^{T}
\end{aligned}
$$

- The variables 'Age', 'Income', 'Outcome' define a basis in $\mathbb{R}^{m}$, and we are interested in dependencies between the variables.


## Correlation

- Correlation is the measure of linear dependency:

$$
\operatorname{Corr}(x, y)=\frac{\operatorname{Cov}(x, y)}{\sqrt{\operatorname{Var}\{x\} \operatorname{Var}\{y\}}}
$$

- If $x=y$, then $\operatorname{Corr}(x, y)=1($ for $\operatorname{Cov}(x, x)=\operatorname{Var}\{x\})$


Uncorrelated
Anticorrelated


$$
\operatorname{Corr}(x, y)=1 \quad \operatorname{Corr}(x, y)=0 \quad \operatorname{Corr}(x, y)=-1
$$

## Correlation matrix

|  | Age | Gender | Income | Outcome | H. owner | C. score |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age | 1,0 | 0,6 | 0,9 | 0,6 | 0,4 | 0,5 |
| Gender | 0,6 | 1,0 | 0,2 | 1,0 | $-0,2$ | $-0,3$ |
| Income | 0,9 | 0,2 | 1,0 | 0,2 | 0,7 | 0,9 |
| Outcome | 0,6 | 1,0 | 0,2 | 1,0 | $-0,2$ | $-0,3$ |
| H. owner | 0,4 | $-0,2$ | 0,7 | $-0,2$ | 1,0 | 0,9 |
| C. score | 0,5 | $-0,3$ | 0,9 | $-0,3$ | 0,9 | 1,0 |

## Principle component analysis

- PCA is a linear transformation of data $y \mapsto K y=x$ :

$$
K y=\left(\begin{array}{ccc}
k_{11} & \ldots & k_{1 m} \\
\vdots & \ddots & \vdots \\
k_{m 1} & \ldots & k_{m m}
\end{array}\right)\left(\begin{array}{c}
y_{1} \\
\vdots \\
y_{m}
\end{array}\right)=\left(\begin{array}{c}
x_{1} \\
\vdots \\
x_{m}
\end{array}\right)=x
$$

- Such that the transformed vectors $x=\left(x_{1}, \ldots, x_{m}\right)$ have uncorrelated coordinates:

$$
\operatorname{Corr}\left(x_{i}, x_{j}\right)=0 \quad \text { for all } \quad i \neq j
$$

- Often most of the variance in the data is accounted by variance in only a few ( $k<m$ ) components (the principal components).


## Correlation $\not \equiv$ dependency

- Let $x \in \mathbb{R}$ and $y$ be defined as:

$$
y=\sin (x)
$$

- Thus, $y$ depends on $x$ by functionally, but

$$
\operatorname{Corr}(x, y)=0
$$

- To see this, recall that correlation represents an average linear trend between $y$ and $x$.
- Generally

$$
\begin{aligned}
& x, y \text { are independent } \\
& x, y \text { are independent }
\end{aligned} \quad \not \operatorname{Corr}(x, y)=0
$$

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## Independence

- Recall that $x$ and $y$ are independent if and only if the conditional probability $P(y \mid x)$ equals to $P(y)$ (marginal):

$$
P(y \mid x)=P(y) \quad \text { or } \quad J(x, y)=Q(x) P(y)
$$

- Dependency is measured by mutual information:

$$
I(x, y):=\mathbb{E}_{J}\left\{\ln \frac{P(y \mid x)}{P(y)}\right\}=\sum_{x, y}\left[\ln \frac{P(y \mid x)}{P(y)}\right] J(x, y) \geq 0
$$

- For dependency in $y=\left(y_{1}, \ldots, y_{m}\right)$ we can consider the divergence:

$$
I\left(y_{1}, \ldots, y_{m}\right)=\sum_{y_{1}, \ldots, y_{m}}\left[\ln \frac{J\left(y_{1}, \ldots, y_{m}\right)}{P\left(y_{1}\right) \otimes \cdots \otimes P\left(y_{m}\right)}\right] J\left(y_{1}, \ldots, y_{m}\right) \geq 0
$$

## Information as distance

- Kullback-Leibler divergence of $Q$ from $P$ in $\mathcal{P}$ :

$$
D_{K L}[P, Q]:=\mathbb{E}_{P}\{\ln P-\ln Q\}=\sum_{\Omega}[\ln P(\omega)-\ln Q(\omega)] P(\omega)
$$

- Surprise associated with observation of event $e \in \Omega$ :

$$
D_{K L}\left[\delta_{e}, Q\right]=\sum_{\Omega}\left[\ln \delta_{e}(\omega)-\ln Q(\omega)\right] \delta_{e}(\omega)=-\ln Q(e)
$$

- Entropy is expected surprise

$$
H[Q]:=\mathbb{E}_{Q}\{-\ln Q\}=-\sum[\ln Q] Q
$$

- Shannon (1948) information is divergence of product of marginals $Q \otimes P$ from joint measure $J$ :

$$
D_{K L}[J, Q \otimes P]=\mathbb{E}_{J}\{\ln J-\ln Q \otimes P\}=: I(x, y)
$$

## Shannon information and entropy

- Shannon (1948) mutual information between $x$ and $y$ :

$$
\begin{aligned}
I(x, y) & =\sum_{X \times Y}\left[\ln \frac{J(x, y)}{Q(x) P(y)}\right] J(x, y) \\
& =\sum_{Y} P(y) \sum_{X}\left[\ln \frac{Q(x \mid y)}{Q(x)}\right] Q(x \mid y) \\
& =H[Q(x)]-H[Q(x \mid y)] \\
& =H[P(y)]-H[P(y \mid x)] \\
& =H[Q(x)]+H[P(y)]-H[J(x, y)]
\end{aligned}
$$

- Shannon information of $x$ is:

$$
I(x, x)=H[Q]
$$

- If $x$ has elementary distribution $\delta_{\omega}(E)$, then:

$$
I(x, x)=H[\delta]=0
$$

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## Blind source separation

- ICA belongs to a class of techniques for blind source separation
- The data $y \in \mathbb{R}^{m}$ that we observe is the result of some unknown transformation $f$ of some unobserved source signals $x \in \mathbb{R}^{n}$ :

$$
y=f(x)
$$

- The goal of BSS is to find the inverse transformation $f^{-1}$ (and hence the sources $\left.x=f^{-1}(y)\right)$ only based on the observed data.
- BSS is possible under some assumptions, such as if $f$ is a linear transformation of independent sources:

$$
y=M x, \quad W \approx M^{-1}, \quad x \approx W y
$$

## Example: The cocktail party problem

- The sources $x=\left(x_{1}, \ldots, x_{m}\right)$ are $m$ people at a party, whose voices are recorded by $n \geq m$ microphones.
- The data $y=\left(y_{1}, \ldots, y_{n}\right)$ are $n$ recordings of mixed signals.




## Independent component analysis

- ICA is a linear transformation of data $y \mapsto W y=x$ :

$$
W y=\left(\begin{array}{ccc}
w_{11} & \ldots & w_{1 m} \\
\vdots & \ddots & \vdots \\
w_{m 1} & \ldots & w_{m m}
\end{array}\right)\left(\begin{array}{c}
y_{1} \\
\vdots \\
y_{m}
\end{array}\right)=\left(\begin{array}{c}
x_{1} \\
\vdots \\
x_{m}
\end{array}\right)=x
$$

- Such that the transformed vectors $x=\left(x_{1}, \ldots, x_{m}\right)$ have independent coordinates:

$$
J\left(x_{1}, \ldots, x_{m}\right)=P\left(x_{1}\right) \otimes \cdots \otimes P\left(x_{m}\right) \quad \text { or } \quad I\left(x_{1}, \ldots, x_{m}\right)=0
$$

- This can be achieved by iterative algorithms that estimate matrix $W$ minimizing $I(W y)$.


## FastICA algorithm

- Recall the Central Limit Theorem, according to which the sum $x_{1}+\cdots+x_{n}$ of $n$ independent random variables with essentially bounded variances converges (in distribution) to a Gaussian random variable.
- Thus, the observed data $y_{i}=w_{i 1} x_{1}+\cdots+w_{i m} x_{m}$ is generally 'more Gaussian' than the independent sources $x_{j}$.
- The non-Gaussianity is measured by neg-entropy, which is approximated by

$$
I\left(y_{i}\right)=\left|\mathbb{E}\left\{G\left(y_{i}\right)\right\}-\mathbb{E}\{G(v)\}\right|^{2}
$$

where $v$ is normal $N(0,1)$ and $G$ are special functions (e.g. $G(u)=(1 / \alpha) \log \cosh (\alpha u)$ or $\left.G(u)=-\exp \left(u^{2} / 2\right)\right)$

- The FastICA algorithm (Hyvärinen \& Oja, 1997) iteratively finds $W \approx M^{-1}$ maximizing $I\left(y_{i}\right)$ for $W y$


## Direct entropy minimization algorithms

- The divergence $I(W y)=I\left(x_{1}, \ldots, x_{n}\right)$ in terms of entropies:

$$
\begin{aligned}
I(W y)= & \sum_{x_{1}, \ldots, x_{n}}\left[\ln \frac{J\left(x_{1}, \ldots, x_{n}\right)}{P\left(x_{1}\right) \otimes \cdots \otimes P\left(x_{n}\right)}\right] J\left(x_{1}, \ldots, x_{n}\right) \\
= & -\left(\sum_{x_{1}}\left[\ln P\left(x_{1}\right)\right] P\left(x_{1}\right)+\cdots+\sum_{x_{n}}\left[\ln P\left(x_{n}\right)\right] P\left(x_{n}\right)\right) \\
& +\sum_{x_{1}, \ldots, x_{n}}\left[\ln J\left(x_{1}, \ldots, x_{n}\right)\right] J\left(x_{1} \ldots, x_{n}\right)=\sum_{i=1}^{n} H\left[P\left(x_{i}\right)\right]-H[J(W y)]
\end{aligned}
$$

- If $W y=x$ is injective, then $H[J(y)]=H[J(W y)]=H[J(x)]$, so that

$$
\min _{W y=x} I(W y) \quad \Longleftrightarrow \quad \max _{W y=x} \sum_{i=1}^{m} H\left[P\left(x_{i}\right)\right]
$$

- RADICAL algorithm does this using Jacobi rotations (Learned-Miller \& Fisher, 2003) and using ordered statistics (Vasicek, 1976) to estimate $H\left[P\left(x_{i}\right)\right]$.





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- Clustering is a partition $X=X_{1} \cup \cdots \cup X_{k}$ of data.
- It is a mapping $f: X \rightarrow Y$ to a set $Y$ of labels (codes):

$$
x \mapsto f(x)=y
$$

- The groups can be based on similaritv.


## Example ( $k$-means)

$Y$ is the set of $k$ points in $(X, d)$, and $f: X \rightarrow Y$ solves:

$$
\min _{f(x)=y} \sum_{i=1}^{k} \sum_{x \inf ^{-1}\left(y_{i}\right)} d\left(x, y_{i}\right)
$$

The new $y_{i} \in Y$ are set to be the centroids $y_{i}^{t+1}=\mathbb{E}\left\{x \inf ^{-1}\left(y_{i}^{t}\right)\right\}$.

## Clustering as source coding

- $f: X \rightarrow Y$ is an encoding, where each $y_{i}$ must have as much information about $x \in X_{i}=f^{-1}\left(y_{i}\right)$ as possible.
- Trivial solution is to use an injective (or uniquely-decodeable) code:

$$
f\left(x_{i}\right)=f\left(x_{j}\right) \quad \Rightarrow \quad x_{i}=x_{j}
$$

- Usually, we want some compression $k=|Y| \ll|X|$ (non-injective $f$ ).
- and preserving as much information as possible:

$$
\max _{f(x)=y} I(x, y)
$$

## Conditional entropy minimization clustering

- for $f(x)=y$

$$
P(y \mid x)=\delta_{f(x)}(y)=\left\{\begin{array}{ll}
1 & \text { if } y=f(x) \\
0 & \text { otherwise }
\end{array} \quad Q(x \mid y)=\frac{Q(x)}{\sum_{x \mathrm{inf}^{-1}(y)} Q(x)}\right.
$$

- Conditional entropy

$$
\begin{aligned}
& H[P(y \mid x)]=0 \\
& \qquad I(x, y)=H[P(y)] \leq \ln |Y|
\end{aligned}
$$

- Maximize $H[P(y)] \leq H[P(x)]$ :

$$
k=|Y| \leq e^{H[P(x)]}
$$

- for $x \inf ^{-1}(y)$
- Conditional entropy $H[Q(x \mid y)] \geq 0$

$$
I(x, y)=H[Q(x)]-H[Q(x \mid y)]
$$

- Minimize $H[Q(x \mid y)]$ :

$$
H[Q(x \mid y)]=\sum^{k} H\left[Q\left(x \inf ^{-1}\left(y_{i}\right)\right)\right]
$$

Detection of HTTP-GET attack
Entropy-based clustering of user online behaviour (Chwalinski, Belavkin, \&


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## Value of information and optimal solutions

- Linear programming problem to find optimal $\hat{P}(y \mid x)=\frac{\hat{J}(x, y)}{Q(x)}$ : minimize $\mathbb{E}_{J}\{d(x, y)\} \quad$ subject to $\quad I(x, y) \leq \lambda$
- The inverse convex programming problem:

$$
\text { minimize } I(x, y) \text { subject to } \quad \mathbb{E}_{J}\{d(x, y)\} \leq v
$$

- Optimal solution for $d(x+a, y+a)=d(x, y)$ (Stratonovich, 1975):

$$
\hat{Q}(x \mid y)=\frac{e^{-\beta d(x, y)}}{\sum_{X} e^{-\beta d(x, y)}}, \quad \beta^{-1}=-\frac{d}{d \lambda} \mathbb{E}_{\hat{\jmath}}\{d\}(\lambda)
$$

- Optimal transformation $x \mapsto y$ given by $\hat{P}(y \mid x)$ is randomized (Belavkin, 2013).


## Geometric value of information

- $\mathbb{E}_{p}\{u\}=\langle u, p\rangle$ expected utility
- $F[p, q]$ information divergence
- Value of information $\lambda$ :

$$
v_{u}(\lambda):=\sup \{\langle u, p\rangle: F[p, q] \leq \lambda\}
$$

- Information of value $v$ :

$$
\lambda_{u}(v):=\inf \{F[p, q]:\langle u, p\rangle \geq v\}=v_{u}^{-1}(v)
$$

- Optimal solutions:
$p(\beta) \in \partial F^{*}[\beta u, q], \quad F[p(\beta), q]=\lambda$
- (Stratonovich, 1965; Belavkin, 2013)


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